

Problem 1. *Answer provided by YANG Mingwei*

The decryption algorithm is $\text{Dec}(k, c) = k^{-1} \cdot c$.

To prove correctness, note that for every $m \in \mathcal{M}$ and $k \in \mathcal{K}$, we have

$$\text{Dec}(k, \text{Enc}(k, m)) = k^{-1} \cdot (k \cdot m) = m.$$

To prove perfect secrecy, as shown in the next problem, it is sufficient to show that for any $m_0, m_1 \in \mathcal{M}$ and $c \in \mathcal{C}$,

$$\Pr [\text{Enc}(k, m_0) = c] = \Pr [\text{Enc}(k, m_1) = c].$$

The left-hand side of the equation equals

$$\Pr [\text{Enc}(k, m_0) = c] = \Pr [k = c \cdot m^{-1}] = \frac{1}{|\mathcal{K}|},$$

so is the right-hand side.

Problem 2.

Perfect secrecy implies perfect indistinguishability. Let M satisfies the uniform distribution over \mathcal{M} . For any $m \in \mathcal{M}$ and $c \in \mathcal{C}$, we have

$$\begin{aligned}\Pr[M = m|C = c] &= \frac{\Pr[M = m] \Pr[C = c|M = m]}{\Pr[C = c]} \\ &= \frac{\Pr[M = m]}{\Pr[C = c]} \cdot \Pr[\text{Enc}(K, m) = c].\end{aligned}\tag{1}$$

Perfect secrecy tells us that $\Pr[M = m|C = c] = \Pr[M = m]$, which implies that $\Pr[\text{Enc}(K, m) = c] = \Pr[C = c]$. Thus for any $m_0, m_1 \in \mathcal{M}$, for any $c \in \mathcal{C}$,

$$\Pr[\text{Enc}(K, m_0) = c] = \Pr[C = c] = \Pr[\text{Enc}(K, m_1) = c].$$

Perfect indistinguishability implies perfect secrecy. Consider any distribution of the message. For arbitrary $c \in \mathcal{C}$ and $m \in \mathcal{M}$

$$\Pr[M = m|C = c] = \frac{\Pr[M = m \wedge C = c]}{\Pr[C = c]}.$$

Its numerator equals $\Pr[M = m] \Pr[\text{Enc}(K, m) = c]$. Its denominator equals

$$\sum_{m' \in \mathcal{M}} \Pr[M = m' \wedge C = c] = \sum_{m' \in \mathcal{M}} \Pr[M = m' | C = c] \Pr[\text{Enc}(K, m') = c].$$

Due to perfect indistinguishability, we know $\Pr[\text{Enc}(K, m') = c] = \Pr[\text{Enc}(K, m) = c]$ for any m' . Therefore, perfect secrecy follows from

$$\begin{aligned}\Pr[M = m|C = c] &= \frac{\Pr[M = m \wedge C = c]}{\Pr[C = c]} \\ &= \frac{\Pr[M = m] \Pr[\text{Enc}(K, m) = c]}{\sum_{m' \in \mathcal{M}} \Pr[M = m'] \Pr[\text{Enc}(K, m') = c]} \\ &= \frac{\Pr[M = m]}{\sum_{m' \in \mathcal{M}} \Pr[M = m']} = \Pr[M = m].\end{aligned}$$

Problem 3. *Answer provided by GUO Chengzhi*

Part B. Proof of $I[X; Y; Z] = I[X; Y] - I[X; Y|Z]$.

$$\begin{aligned}
 I[X; Y; Z] &= H[X] + H[Y] + H[Z] - H[X, Y] - H[X, Z] - H[Y, Z] + H[X, Y, Z] \\
 &= I[X; Y] + H[Z] - H[X, Z] - H[Y, Z] + H[X, Y, Z] \\
 &= I[X; Y] - (H[X, Z] - H[Z]) - (H[Y, Z] - H[Z]) - (-H[X, Y, Z] + H[Z]) \\
 &= I[X; Y] - H[X|Z] - H[Y|Z] + H[X, Y|Z] \\
 &= I[X; Y] - I[X; Y|Z]
 \end{aligned}$$

Proof of $H[Z] \geq I[X; Y; Z] \geq -H[Z]$.

$$\begin{aligned}
 I[X; Y; Z] &= H[X] + H[Y] + H[Z] - H[X, Y] - H[X, Z] - H[Y, Z] + H[X, Y, Z] \\
 &= (H[X] - H[Z, X]) + (H[Y] - H[X, Y]) + (H[Z] - H[Y, Z]) + H[X, Y, Z] \\
 &= -H[Z|X] - H[X|Y] - H[Y|Z] + H[X, Y, Z] \\
 &= -H[Z|X] - H[X|Y] - H[Y|Z] + H[Z] + H[Y|Z] + H[X|Y, Z] \\
 &= -H[Z|X] - H[X|Y] + H[Z] + H[X|Y, Z]
 \end{aligned}$$

Since $H[X|Y, Z] - H[X|Y] \leq 0$, $-H[Z|X] \leq 0$, we have

$$I[X; Y; Z] = -H[Z|X] - H[X|Y] + H[Z] + H[X|Y, Z] \leq H[Z]$$

On the other hand, we have

$$\begin{aligned}
 H[Z] - H[Z|X] &\geq 0 \\
 H[X|Y, Z] - H[X|Y] &= -I[X; Z|Y] \geq -H[Z|Y] \geq -H[Z],
 \end{aligned}$$

so we get

$$I[X; Y; Z] \geq -H[Z]$$

An example where $I[X; Y; Z] = H[Z] > 0$. Assume Z is taken uniformly from $\{0, 1\}$ and $X = Y = Z$, then $I[X; Y; Z] = H[Z] = 1$.

An example where $I[X; Y; Z] = -H[Z] < 0$. Assume X, Y are taken uniformly from $\{0, 1\}$ and $Z = X \oplus Y$, then $I[X; Y; Z] = -H[Z] = -1$.

Part C. Let random variable M, K, C denote the message, key, ciphertext during an encryption. The distribution of K is specified by **Gen**. The distribution of M will be specified later. First we have

$$I[C; M|K] = H[C|K] + H[M|K] - H[C, M|K].$$

Since M and K are independent, we have

$$H[M|K] = H[M].$$

Since M can be determined by C, K , we have $H[M|C, K] = 0$. So

$$H[C, M|K] = H[C|K] + H[M|C, K] = H[C|K].$$

Therefore

$$I[C; M|K] = H[M]$$

By perfect secrecy, we have $I[C; M] = 0$, so by Part B we have

$$\begin{aligned} I[C; M; K] &= I[C; M] - I[C; M|K] = -H[M] \\ H[K] &\geq -H[M] \geq -H[K] \end{aligned}$$

Let M satisfy the uniform distribution over \mathcal{M} , then $H[M] = \log |\mathcal{M}|$. By inequality above we have $H[K] \geq \log |\mathcal{M}|$.

Problem 4.

Part A. $|\mathcal{K}| \geq |\mathcal{M}|$

Proof. Suppose not. Take arbitrary c_0 . Since $|\{\text{Dec}(k, c_0) : k \in \mathcal{K}\}| \leq |\mathcal{K}| < |\mathcal{M}|$, there must exist m such that $\Pr[M = m | C = c_0] = 0$. Then, for any distribution where $\Pr[M = m] > 1 - \varepsilon$, the condition $|\Pr[M = m | C = c_0] - \Pr[M = m]| \leq 1 - \varepsilon$ cannot hold. So $|\mathcal{K}| \geq |\mathcal{M}|$.

Part B. $|\mathcal{K}| \geq (1 - \varepsilon)|\mathcal{M}|$

Fix m_0 to be any message, sample m_1 uniformly at random. Consider the following distinguisher D .

$$D(c) \text{ outputs } 1 \iff \exists k \in \mathcal{K} \text{ such that } \text{Dec}(k, c) = m_1.$$

Since the scheme is perfectly correct, the distinguisher always outputs 1 when $b = 1$.

When $b = 0$, let $c = \text{Enc}(K, m_0)$ be the ciphertext. The distinguisher outputs 1 if and only if there exists k such that $\text{Dec}(k, c) = m_1$. In other words, the distinguisher outputs 1 if and only if

$$m_1 \in \left\{ \text{Dec}(k, c) \mid k \in \mathcal{K} \right\}.$$

Note that m_1 is independent from k, c , so the probability m_1 falls in the above set is at most $|\mathcal{K}|/|\mathcal{M}|$.

$$\begin{aligned} & \Pr_{\substack{k \leftarrow \text{Gen} \\ b \leftarrow \{0,1\}}} \left[D(\text{Enc}(K, m_b)) = b \right] \\ &= \frac{1}{2} \Pr_{k \leftarrow \text{Gen}} \left[D(\text{Enc}(K, m_1)) = 1 \right] + \frac{1}{2} \Pr_{k \leftarrow \text{Gen}} \left[D(\text{Enc}(K, m_0)) \neq 1 \right] \\ &\geq \frac{1}{2} \left(1 + 1 - \frac{|\mathcal{K}|}{|\mathcal{M}|} \right). \end{aligned}$$

Part C. $|\mathcal{K}| \geq (1 - \varepsilon)|\mathcal{M}|$

Proof. (The proof assumes algorithms Enc, Dec to be deterministic. While it can be easily generated to the case where Enc, Dec may be probabilistic.) Suppose not. Take the uniform distribution over \mathcal{M} . Fix arbitrary $c_0 \in \mathcal{C}$.

Since for all m ,

$$\Pr[M = m | C = c_0] = \Pr[M = m] = \frac{1}{|\mathcal{M}|}$$

we can compute

$$\begin{aligned} \Pr_{k \leftarrow \text{Gen}} [\text{Enc}(m, k) = c_0] &= \Pr[C = c_0 | M = m] \\ &= \Pr[M = m | C = c_0] \cdot \frac{\Pr[C = c_0]}{\Pr[M = m]} \\ &= \Pr[C = c_0] \end{aligned}$$

Let $S_{c_0} = \{\text{Dec}(k, c_0) : k \in \mathcal{K}\}$. We know

$$|S_{c_0}| \leq |\mathcal{K}| < (1 - \varepsilon)|\mathcal{M}|$$

For $m \notin S_{c_0}$, if $\text{Enc}(k, m) = c_0$, then $\text{Dec}(k, \text{Enc}(k, m)) \neq m$. Thus,

$$\begin{aligned}
& \Pr[\text{Dec}(K, \text{Enc}(K, M)) \neq M] \\
& \geq \Pr[M \notin S_{\text{Enc}(K, M)}] \\
& = \sum_{c_0 \in \mathcal{C}} \Pr[M \notin S_{c_0}, C = c_0] \\
& = \sum_{c_0 \in \mathcal{C}} \sum_{m \notin S_{c_0}} \Pr[C = c_0, M = m] \\
& = \sum_{c_0 \in \mathcal{C}} \sum_{m \notin S_{c_0}} \Pr[C = c_0 | M = m] \Pr[M = m] \\
& = \frac{1}{|\mathcal{M}|} \sum_{c_0 \in \mathcal{C}} \sum_{m \notin S_{c_0}} \Pr_{k \leftarrow \text{Gen}} [\text{Enc}(m, k) = c_0] \\
& = \frac{1}{|\mathcal{M}|} \sum_{c_0 \in \mathcal{C}} \sum_{m \notin S_{c_0}} \Pr[C = c_0] \\
& = \frac{1}{|\mathcal{M}|} \sum_{c_0 \in \mathcal{C}} \Pr[C = c_0] (|\mathcal{M}| - |S_{c_0}|) \\
& > \frac{\varepsilon |\mathcal{M}|}{|\mathcal{M}|} \sum_{c_0 \in \mathcal{C}} \Pr[C = c_0] \\
& = \varepsilon.
\end{aligned}$$

This also implies $\Pr[\text{Dec}(K, \text{Enc}(K, M)) = M] < 1 - \varepsilon$. But on the other hand,

$$\begin{aligned}
& \Pr[\text{Dec}(K, \text{Enc}(K, M)) = M] \\
& = \sum_{m \in \mathcal{M}} \Pr[\text{Dec}(K, \text{Enc}(K, M)) = M | M = m] \Pr[M = m] \\
& = \sum_{m \in \mathcal{M}} \Pr_{k \leftarrow \text{Gen}} [\text{Dec}(k, \text{Enc}(k, m)) = m | M = m] \Pr[M = m] \\
& \geq (1 - \varepsilon) \sum_{m \in \mathcal{M}} \Pr[M = m] \\
& = 1 - \varepsilon.
\end{aligned}$$

This is a contradiction.

An Alternative Proof. Let random variable M be uniformly distributed over \mathcal{M} . Since the scheme is perfectly secure, by Problem 2, we know that for any $m_0, m_1 \in \mathcal{M}$, it holds that

$$\forall c \in \mathcal{C}, \Pr[\text{Enc}(K, m_0) = c] = \Pr[\text{Enc}(K, m_1) = c]$$

which implies that $\text{Enc}(K, M)$ and M are independent. Define oracle $\text{Eve}(m, c)$ for any $m \in \mathcal{M}, c \in \mathcal{C}$ as

$$\text{Eve}(m, c) = [\exists k \in \mathcal{K}, \text{Dec}(k, c) = m].$$

On one hand, we know that

$$\begin{aligned}
\Pr[\text{Eve}(M, \text{Enc}(K, M)) = 1] &= \Pr[\exists k \in \mathcal{K}, \text{Dec}(k, \text{Enc}(K, M)) = M] \\
&\geq \Pr[\text{Dec}(K, \text{Enc}(K, M)) = M] \\
&\geq 1 - \varepsilon
\end{aligned}$$

On the other hand,

$$\Pr[\text{Eve}(M, \text{Enc}(K, M)) = 1] \leq \sum_{k \in \mathcal{K}} \Pr[\text{Dec}(k, \text{Enc}(K, M)) = M] \leq \sum_{k \in \mathcal{K}} \frac{1}{|\mathcal{M}|} = \frac{|\mathcal{K}|}{|\mathcal{M}|}$$

where the second inequality is because M and $\text{Enc}(K, M)$ are independent, and M is uniformly distributed. Therefore, it follows that

$$1 - \varepsilon \leq \frac{|\mathcal{K}|}{|\mathcal{M}|}$$

by which we obtain $|\mathcal{K}| \geq (1 - \varepsilon)|\mathcal{M}|$.